

New environmental dependent modelling with Gaussian particle filtering based implementation for ground vehicle tracking

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Outline of the presentation

- Overview of the intelligent transportation system
- Traditional methods for the ground vehicle state estimation
- > The developed method for tracking the ground vehicle:
- Model
- Algorithm
- Simulation studies
- Conclusions and Future plans

Overview of the intelligent transport system



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Reviews of the traffic participants state estimation methods

Filtering based algorithm:



Domain knowledge aided:

Physical constraints (road network)



M. Ulmke and W. Koch, "Road-map assisted ground moving target tracking," *Aerospace and Electronic Systems, IEEE Transactions on*, vol. 42, no. 4, pp. 1264–1274, 2006.

H. Oh, H. Shin, S. Kim, A. Tsourdos, and B. White, "Airborne behaviour monitoring using Gaussian processes with map information," *IET Radar, Sonar and Navigation*, vol. 7, no. 4, pp. 393–400, 2013.



Context information



T. Gindele, S. Brechtel, and R. Dillmann, "A probabilistic model for estimation driver behaviors and vehicle trajectory in traffic environment," in 13th International IEEE Annual Conference on Intelligent Transportation Systems, Madeira Island, Portugal, 2010.

Liu, C. and Chen, B. L. W., "Road network based vehicle navigation using an improved IMM particle filter," 2013 IFAC Intelligent Autonomous Vehicles Symposium, Gold Coast, Australia (2013).



Outline of the proposed algorithm



Non-Markov jump modelling for incorporating context information



 x_t : state m_t : movement mode C_t : context

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State model refinement

A general state model in the non-Markov jump model system: $x_{t+1} = F_{m_t}(x_t) + Gw_t \rightarrow p(x_{t+1}|x_t, m_t, Y_{t-1})$

X. Li and P. Jilkov, "Survey of maneuvering target tracking. Part I: Dynamic models", IEEE Trans. on Aerospace and Electronic Systems, vol. 39, no. 4, pp. 1333–1364, 2003.

Two-step refinement:

I. Social force:



Originally developed for modelling the human movements

Helbing, D., Farkas, I., Viscek, T.: 'Simulating dynamic features of escape panic', Nature, 2000, 407, pp. 487–490.

Extended for modelling the traffic participants in a shared space (including vehicles):

B. Anvaria, M. Bella, A. Sivakumara, W. Ochienga, " Modelling Shared Space Users Via Rule-based Social Force Model ", Transportation Research Part C: Emerging Technologies, Volume 51, 2015, Pages 83–103



Forces:

$$f_{i,o}^{repulsive} = A \cdot \exp(\frac{-d_{i,o}}{B}) \boldsymbol{n}_{io}$$
$$f_{i,c}^{attractive} = A' \cdot (1 - \exp(\frac{-d_{i,c}}{B'})) \boldsymbol{n}_{io}$$

Refined model:

$$x_{t+1} = F_{m_t}(x_t) + I(f_t) + Gw_t \to p_f(x_{t+1}|x_t, m_t, Y_{t-1})$$



II. Constraints:

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Algorithms development—Bayesian inference



 $p(\mathbf{x}_{t}, m_{t} | \mathbf{Y}_{t}) \propto p(m_{t} | \mathbf{Y}_{t-1}) p(\mathbf{y}_{t} | \mathbf{x}_{t}, m_{t}) p(\mathbf{x}_{t} | m_{t}, \mathbf{Y}_{t-1})$ •

Probability inference flowchart:



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Algorithms development—Generic particle filtering based implementation

Initially, we have a set of particles $\{\mathbf{x}_{t-1}^{r,i}, \mathbf{u}_{t}^{r,i}\}_{i=1,\dots,N}$ to approximate $p(m_{t-1} = r, \mathbf{x}_{t-1} | Y_{t-1})$ for every mode r:

> Mode mixing:

 $p(m_t = m | \boldsymbol{Y}_{t-1}) \approx \sum_{r \in M} \sum_{i=1}^{N} p(m_t = m | \boldsymbol{C}(\boldsymbol{x}_{t-1}^{r,i}, environment), m_{t-1} = r) u_t^{r,i} \equiv \Lambda_{t-1}^{u}$

> Interacting:

$$p(\mathbf{x}_{t-1}|m_t = m, \mathbf{Y}_{t-1}) \approx \frac{\sum_{r \in M} \sum_{i=1}^{N} p(m_t = m | m_{t-1} = r, \mathbf{x}_{t-1}^{r,i}) u_{t-1}^{r,i} \delta(x_{t-1} - x_{t-1}^{r,i})}{\Lambda_{t-1}^{u}}$$

N particles $\{\overline{x}_{t-1}^{m,i}, \overline{u}_{t-1}^{m,i}\}_{i=1,...,N}$ are resampled to solve the exponential increasing of particles number

> Evolution:

Sampling particles
$$\{x_t^{m,i}\}_{i=1,\dots,N}$$
 from $p_f^C(x_t|m_t = m, \overline{x}_{t-1}^{m,i}, Y_{t-1})$

$$p(\mathbf{x}_t | m_t = m, \mathbf{Y}_{t-1}) \approx \bar{u}_{t-1}^{m,i} \delta(\mathbf{x}_{t-1} - \mathbf{x}_{t-1}^{m,i})$$

> Correction:

$$p(\mathbf{x}_{t}, m_{t} = m | \mathbf{Y}_{t-1}) \propto \sum_{i=1}^{N} p(m_{t} = m | \mathbf{Y}_{t-1}) p(\mathbf{y}_{t} | \mathbf{x}_{t}^{m,i}, m_{t} = m) \bar{u}_{t-1}^{m,i} \delta(\mathbf{x}_{t} - \mathbf{x}_{t}^{m,i})$$

H. Blom and E. Bloem, "Exact Bayesian and particle filtering of stochastic hybrid systems", IEEE Transaction on Aerospace and Electronic Systems, Vol.43, No.1, Pages 55-70, 2007.

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Algorithms development—Gaussian particle filtering based implementation

Initially, we have a set of particles $\{\mathbf{x}_{t-1}^{r,i}, \mathbf{u}_{t}^{r,i}\}_{i=1,\dots,N}$ to approximate $p(m_{t-1} = r, \mathbf{x}_{t-1}|Y_{t-1})$ for every mode r

Mode mixing:

 $p(m_t = m | \mathbf{Y}_{t-1}) \approx \sum_{r \in M} \sum_{i=1}^{N} p(m_t = m | \mathbf{x}_{t-1}^{r,i}, m_{t-1} = r) u_t^{r,i} \equiv \Lambda_{t-1}^{u}$

Interacting:

$$p(\mathbf{x}_{t-1}|m_t = m, \mathbf{Y}_{t-1}) \approx \frac{\sum_{r \in M} \sum_{i=1}^{N} p(m_t = m | m_{t-1} = r, \mathbf{x}_{t-1}^{r,i}) u_{t-1}^{r,i} \delta(x_{t-1} - x_{t-1}^{r,i})}{\Lambda_{t-1}^u}$$

N particles $\{\overline{x}_{t-1}^{m,i}, \overline{u}_{t-1}^{m,i}\}_{i=1,\dots,N}$ are resampled to solve the exponential increasing of particles number

> Evolution and correction:

Initialization of mean and covariance : $\hat{x}_{t-1|t-1} = \sum_{i=1}^{N} \bar{u}_{t-1}^{m,i} \cdot \bar{x}_{t-1}^{m,i}$ $\hat{P}_{t-1|t-1} = \sum_{i=1}^{N} \bar{u}_{t-1}^{m,i} \cdot (\bar{x}_{t-1}^{m,i} - \hat{x}_{t-1|t-1}) \cdot (\bar{x}_{t-1}^{m,i} - \hat{x}_{t-1|t-1})^T$

> A Gaussian distribution $N_c(x_t | \hat{x}_{t|t}, \hat{P}_{t|t})$ is obtained using <u>truncated unscented Kalman filter</u>

> New particles $\{x_t^{m,i}\}_{i=1,...,N}$ are sampled from $N_c(x_t | \hat{x}_{t|t}, \hat{P}_{t|t})$ for approximating $p(x_t | m_t = m, Y_t)$ as:

$$p(\mathbf{x}_{t}, m_{t} = m | \mathbf{Y}_{t}) \propto \sum_{i=1}^{N} u_{t}^{m,i} \delta(\mathbf{x}_{t} - \mathbf{x}_{t}^{m,i}), \quad with \quad u_{t}^{m,i} = \frac{p(\mathbf{y}_{t} | \mathbf{x}_{t}, m_{t} = m) N(\mathbf{x}_{t}^{m,i} | \hat{\mathbf{x}}_{t|t-1}, \hat{P}_{t|t-1})}{N(\mathbf{x}_{t}^{m,i} | \hat{\mathbf{x}}_{t|t}, P_{t})}$$

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Truncated unscented Kalman filter

Based on the initial mean \overline{x} and covariance *P* of x_{t-1} related to $p(x_{t-1}|m_t, Y_{t-1})$

Obtaining σ -points:

$$\begin{aligned} \mathbf{x}_0 &= \overline{\mathbf{x}} \quad u_0 = \frac{\kappa}{n_{\chi} + \kappa} \quad \mathbf{y}_0 = h(\overline{\mathbf{x}}, \mathbf{e}_0) \\ \mathbf{x}_i &= \overline{\mathbf{x}} + (\sqrt{(n_{\chi} + \kappa)P})_i \quad u_i = \frac{1}{2(n_{\chi} + \kappa)} \quad \mathbf{y}_i = h(\mathbf{x}_i, \mathbf{e}_i) \\ \mathbf{x}_{n_{\chi} + i} &= \overline{\mathbf{x}} - (\sqrt{(n_{\chi} + \kappa)P})_i \quad u_{n_{\chi} + i} = u_i \quad \mathbf{y}_{n_{\chi} + i} = h(\mathbf{x}_{n_{\chi} + i}, \mathbf{e}_{n_{\chi} + i}) \end{aligned}$$

Mean and covariance updating:

$$P_{\mathbf{y}_{t}|\mathbf{y}_{t}} = \sum_{j=0}^{2n_{\chi}} u_{j} (\mathbf{y}_{j} - \widehat{\mathbf{y}}) (\mathbf{y}_{j} - \widehat{\mathbf{y}})^{T} \quad \text{where } \widehat{\mathbf{y}} = \sum_{j=0}^{2n_{\chi}} u_{j} \mathbf{y}_{j}$$

$$P_{\mathbf{x}_{t}|\mathbf{y}_{t}} = \sum_{j=0}^{2n_{\chi}} u_{j} (\mathbf{x}_{j} - \overline{\mathbf{w}}) (\mathbf{y}_{j} - \widehat{\mathbf{y}})^{T}$$

$$K = P_{\mathbf{x}_{t}|\mathbf{y}_{t}} P_{\mathbf{y}_{t}|\mathbf{y}_{t}}^{-1}$$

$$\overline{\mathbf{x}'} = \overline{\mathbf{x}} + K(\mathbf{y}_{t} - \widehat{\mathbf{y}})$$

$$P' = P - K P_{\mathbf{y}_{t}|\mathbf{y}_{t}} K^{T}$$

Constraint information incorporation:

M samples $\{x_i\}_{i=1,...,M}$ are obtained from $N(\mathbf{x}_t \mid \overline{x'}, P')$, new mean and covariance are obtained after considering the constraint:

$$\overline{\mathbf{x}'}_{C} = \frac{1}{N} \sum_{j=1}^{M'} \mathbf{x}_{j}$$

$$P'_{C} = \frac{1}{N} \sum_{j=1}^{M'} (\mathbf{x}_{j} - \overline{\mathbf{x}'}_{C}) (\mathbf{x}_{j} - \overline{\mathbf{x}'}_{C})^{T}$$

where $\{x_{j}\}_{i=1,...,M} \in C$

Samples are obtained from $N(\mathbf{x}_t | \overline{\mathbf{x}'}_c, \mathbf{P'}_c)$ incorporating the measurement and constraint information

O. Straka, J. Dunik and M. Simandl, "Truncated nonlinear filters for state estimation with nonlieanr inequality constraints", Automatica, Vol.48, No.2, Pages 273-286, 2012.







Simulation studies--scenario

Simulated vehicles moving scenario and trajectories:





Vehicles manoeuvring behaviours:



Simulation studies--models

State models:

$$\begin{array}{lll} \text{Constant velocity model:} & \begin{pmatrix} x_k \\ y_k \\ \dot{x}_k \\ \dot{y}_k \end{pmatrix} = \begin{pmatrix} 1 \ 0 \ T \ 0 \\ 0 \ 1 \ 0 \ T \\ \dot{y}_{k-1} \\ \dot{y}_{k-1} \\ \dot{y}_{k-1} \\ \dot{y}_{k-1} \\ \dot{y}_{k-1} \end{pmatrix} + \begin{pmatrix} T^2/2 & 0 \\ 0 & T^2/2 \\ T & 0 \\ 0 & T \end{pmatrix} \begin{pmatrix} n_x \\ n_y \end{pmatrix} & N(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 0.1^2(m/s^2)^2, 0 \\ 0, 0.1^2(m/s^2)^2 \end{pmatrix} \\ N(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 2^2(m/s^2)^2, 0 \\ 0, 1^2(m/s^2)^2 \end{pmatrix} \\ N(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 2^2(m/s^2)^2, 0 \\ 0, 1^2(m/s^2)^2 \end{pmatrix} \\ N(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 2^2(m/s^2)^2, 0 \\ 0, 1^2(m/s^2)^2 \end{pmatrix} \\ \begin{pmatrix} x_k \\ y_k \\ \dot{y}_k \\ \dot{y}_k \\ w_k \end{pmatrix} = \begin{pmatrix} 1 & \frac{\sin wT}{w} & 0 & -\frac{1-\cos wT}{w} & 0 \\ 0 & \cos wT & 0 & -\sin wT & 0 \\ 0 & \sin wT & 0 & \cos wT & 0 \\ 0 & \sin wT & 0 & \cos wT & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} x_k \\ y_k \\ \dot{y}_k \\ w_k \end{pmatrix} + \begin{pmatrix} T^2/2 & 0 & 0 \\ T & 0 & 0 \\ 0 & T^2/2 & 0 \\ 0 & T & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} n_x \\ n_y \\ n_w \end{pmatrix})$$

Each model is refined by both force (from the road boundaries and another vehicle) and constraint (road)

Model transitions:

Model transition probabilities are defined in a context dependent way, e.g.

 $p(m_t = HICV | m_t = LICV) = exp(-a \cdot d_{vehicle}),$ $p(m_t = CT | m_t = LICV) = exp(-a' \cdot d_{turn}),$ a = 0.1, a' = 0.15: empirical set parameters

 $d_{vehicle}$: distance to the front vehicle d_{turn} : distance to the turn



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Measurements:

A sensor positioned at [200,30] (m) is applied to measure the range r_t and bearing angle θ_t of a particular vehicle with:

$$\hat{y}_{t}^{t} = \begin{pmatrix} \sqrt{(x_{s} - x_{t})^{2} + (y_{s} - y_{t})^{2}} \\ \tan^{-1} \frac{y_{s} - y_{t}}{x_{s} - x_{t}} \end{pmatrix} + {n_{r} \choose n_{\theta}}, {n_{r} \choose n_{\theta}} \sim N({0 \choose 0}, {5^{2}(m^{2}), 0 \choose 0, 0.02^{2}(rad^{2})})$$

Simulation studies—manoeuvring type determination



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Simulation studies—modelling comparisons

- > Comparisons between different modelling approaches are made from 100 Monte-Carlo simulations
- > Gaussian particle filtering based approach (with the same particle number) is applied for every model



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Simulation studies—algorithms comparison

Under the same model, comparisons have been made with the generic particle filtering based and Gaussian particle based approaches



AVE_RMSEs (s) for different implementation algorithms

	Generic particle filtering	Gaussian particle filtering
Vehicle 1	2.77	1.51
Vehicle 2	2.65	2.10

Computational costs for trajectories tracking (per sample):

- Generic particle filtering: 0.12 (s)
- Gaussian particle filtering: 0.03 (s)

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Conclusions and future works

> In this work, we have developed a new traffic participants state estimation method:

Model aspect:

- Non-Markov jump model is applied to model the transition probabilities based on the context information
- For a particular model, both the force and constraint are applied for refining the model

Algorithm aspect:

- Bayesian inference algorithm is developed for the model
- Gaussian particle filtering based algorithm is applied for implementation
- > Future work:
- Data associations (miss detection/false alarms)
- System parameters calibration (machine learning)



Questions?