







Multi-Polarization SAR Change Detection: Unstructured Versus Structured GLRT

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- Change detection is the capability to identify temporal changes within a given scene starting from a pair of co-registered images representing the area of interest;
- Two main approaches exists: incoherent and coherent;
- We introduce a structured approach to derive the Generalized Likelihood Ratio Test;
- The block diagonal structure of the polarimetric covariance matrix is exploited to achieve our goal, moreover the derived rule show Constant False Alarm Rate (CFAR) behavior.





A multipolarization SAR sensor measures for each pixel of the image under test $N \in \{2,3\}$ complex returns, collected from different polarimetric channels.

The N returns from the same pixel are stacked to form the vector $\mathbf{X}(I, m)$, where $I = 1, \dots, L$ and $m = 1, \dots, M$.











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- We focus on the problem of detecting the presence of possible changes in a rectangular neighbourhood A, with size $K = W_1 \times W_2 \ge N$, of a given pixel;





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- We focus on the problem of detecting the presence of possible changes in a rectangular neighbourhood A, with size $K = W_1 \times W_2 \ge N$, of a given pixel;
- We denote by \mathbf{R}_X (\mathbf{R}_Y) the matrix whose columns are the vectors of the polarimetric returns from the pixels of **X** (**Y**) which fall in the region \mathcal{A} and $\mathbf{S}_X = \mathbf{R}_X \mathbf{R}_X^{\dagger}$ ($\mathbf{S}_Y = \mathbf{R}_Y \mathbf{R}_Y^{\dagger}$).









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- The columns of \mathbf{R}_{x} (\mathbf{R}_{Y}) are assumed statistically i.i.d. random vectors drawn from a complex circular zero-mean Gaussian distribution with positive definite covariance matrix $\mathbf{\Sigma}_{X}$ ($\mathbf{\Sigma}_{Y}$), complying with the structure introduced in [Novak 1990]

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where

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In this way, the change detection problem in the region A can be formulated in terms of the following binary hypothesis test

$$\begin{cases} H_0: \quad \mathbf{\Sigma}_X = \mathbf{\Sigma}_Y \\ H_1: \quad \mathbf{\Sigma}_X \neq \mathbf{\Sigma}_Y \end{cases}$$





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Unstructured and Structured GLRT









The Unstructured GLRT does not exploit the special structure of Σ_X and Σ_Y and was derived in [Novak 2005]

$$\frac{\det^2(\mathbf{S}_X + \mathbf{S}_Y)}{\det(\mathbf{S}_X)\det(\mathbf{S}_Y)} \stackrel{H_1}{\underset{H_0}{\geq}} T_U,$$









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In our approach we consider the special structure of the covariance \boldsymbol{S}_X and \boldsymbol{S}_Y

$$\mathbf{S}_{X} = \begin{bmatrix} \mathbf{S}_{X,1} & \mathbf{S}_{X,2} \\ \mathbf{S}_{X,2}^{\dagger} & \widehat{\sigma^{2}}_{X,1} \end{bmatrix} \qquad \qquad \mathbf{S}_{Y} = \begin{bmatrix} \mathbf{S}_{Y,1} & \mathbf{S}_{Y,2} \\ \mathbf{S}_{Y,2}^{\dagger} & \widehat{\sigma^{2}}_{Y,1} \end{bmatrix},$$







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- The structured GLRT is the decision rule

$$\frac{\sum_{X,1}\sum_{Y,1,\sigma_{X,1}^2,\sigma_{Y,1}^2} f_{\mathbf{R}_X,\mathbf{R}_Y}(\mathbf{R}_X,\mathbf{R}_Y|H_1,\boldsymbol{\Sigma}_{X,1}\boldsymbol{\Sigma}_{Y,1},\sigma_{X,1}^2,\sigma_{Y,1}^2)}{\sum_{X,1,\sigma_{X,1}^2} f_{\mathbf{R}_X,\mathbf{R}_Y}(\mathbf{R}_X,\mathbf{R}_Y|H_0,\boldsymbol{\Sigma}_{X,1},\sigma_{X,1}^2)} \stackrel{H_1}{\underset{H_0}{\overset{>}{\underset{K_{Y,1}}{\underset{K_{Y,1}}{\overset{>}{\underset{K_{Y,1}}{\overset{>}{\underset{K_{Y,1}}{\overset{>}{\underset{K_{Y,1}}{\underset{K_$$









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By replacing the unknown parameters in the likelihood ratio with their maximum likelihood estimates, under each hypothesis, we obtain the structured GLRT.







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Unstructured and Structured GLRT

 Hence, performing the maximizations over the parameters we can obtain the structured GLRT

$$\frac{\det^{2\mathcal{K}}(\mathbf{S}_{X,1}+\mathbf{S}_{Y,1})}{\det^{\mathcal{K}}(\mathbf{S}_{X,1})\det^{\mathcal{K}}(\mathbf{S}_{Y,1})}\frac{\left(\widehat{\sigma^{2}}_{X,1}+\widehat{\sigma^{2}}_{Y,1}\right)^{2\mathcal{K}}}{\left(\widehat{\sigma^{2}}_{X,1}\widehat{\sigma^{2}}_{Y,1}\right)^{\mathcal{K}}} \stackrel{H_{1}}{\underset{H_{0}}{\overset{<}{\underset{K}}}} T_{\mathcal{S},1},$$

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$$\frac{\det^{2K}(\mathbf{S}_{X,1}+\mathbf{S}_{Y,1})}{\det^{K}(\mathbf{S}_{X,1})\det^{K}(\mathbf{S}_{Y,1})}\frac{\left(\widehat{\sigma^{2}}_{X,1}+\widehat{\sigma^{2}}_{Y,1}\right)^{2K}}{\left(\widehat{\sigma^{2}}_{X,1}\widehat{\sigma^{2}}_{Y,1}\right)^{K}} \stackrel{H_{1}}{\underset{H_{0}}{\overset{2}{\underset{K}{\longrightarrow}}} T_{\mathcal{S},1},$$

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Finally, after a monotonic transformation, we get the following equivalent form of the GLRT

$$\frac{\det^2(\mathbf{S}_{X,1}+\mathbf{S}_{Y,1})}{\det(\mathbf{S}_{X,1})\det(\mathbf{S}_{Y,1})}\frac{\left(\widehat{\sigma^2}_{X,1}+\widehat{\sigma^2}_{Y,1}\right)^2}{\widehat{\sigma^2}_{X,1}\widehat{\sigma^2}_{Y,1}} \stackrel{H_1}{\underset{H_0}{\overset{>}{\rightarrow}} T_S,$$

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with T_S the modified detection threshold.

It can be proved that this decision rule ensures the CFAR property with respect to both $\Sigma_{X,1}$ and $\sigma_{X,1}^2$.









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 The standard ROCs are computed for the unstructured and structured GLRTs and compared with the benchmark performance of the optimum Neyman-Pearson detector;







- The standard ROCs are computed for the unstructured and structured GLRTs and compared with the benchmark performance of the optimum Neyman-Pearson detector;
- The optimum receiver assumes that the actual covariance matrices are known, and can be expressed as:

$$tr\left[\left(\boldsymbol{\Sigma}_{X}^{-1}-\boldsymbol{\Sigma}_{Y}^{-1}\right)\boldsymbol{S}_{Y}\right] \begin{array}{c} H_{1} \\ \stackrel{>}{\leq} \\ H_{0} \end{array} T,$$

which resorting to the special structure of $\pmb{\Sigma}_X$ and $\pmb{\Sigma}_Y$ leads to

$$tr\left[\left(\boldsymbol{\Sigma}_{X,1}^{-1}-\boldsymbol{\Sigma}_{Y,1}^{-1}\right)\boldsymbol{S}_{Y,1}+\left(\frac{1}{\sigma_{X,1}^2}+\frac{1}{\sigma_{X,2}^2}\right)\hat{\sigma}_{Y,1}^2\right] \begin{array}{c} H_1\\ \stackrel{>}{<}\\ H_0\\ H_0 \end{array} T.$$





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In order to set the detection threshold, Monte Carlo simulations are used assuming $100/P_{fa}$ independent runs. Additionally, 10^5 independent trials are exploited to estimate P_d ;





- In order to set the detection threshold, Monte Carlo simulations are used assuming $100/P_{fa}$ independent runs. Additionally, 10^5 independent trials are exploited to estimate P_d ;
- The theoretical covariance matrices considered to estimate the P_d are:

$$\mathbf{\Sigma}_X = \begin{pmatrix} 1 & 0.5 & 0 \\ 0.5 & 1 & 0 \\ 0 & 0 & 0.2 \end{pmatrix} \qquad \mathbf{\Sigma}_Y = 2\mathbf{\Sigma}_X,$$

while $\Sigma_Y = \Sigma_X$ was considered to estimate the P_{fa} .





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Performance Assessment



 P_d versus P_{fa} for W = [3, 5, 7].







- The analysis is performed using real X-band data available in the Coherent Change Detection Challenge dataset acquired by the Air Force Research Laboratory (AFRL).
 The dataset has been acquired using a coherent receiver with 640 MHz bandwidth and dual-polarized mode with a range and cross-range resolution of 0.3 m;
- The selected area of interest is a sub-image of 1000 × 1000 pixels (i.e., L = M = 1000) and is composed of several parking lots which are occupied by numerous parked, (i.e., stationary) vehicles.





For this particular scenario the changes between the reference and test images (denoted by X and Y respectively), occurred during the time interval between the two acquisitions can be distinguished in two cases:

- a vehicle is present in X but is not present in Y, this case is referred as departure;
- a vehicle is not present in X but is present in Y, this event is referred as arrival.

Using the cases defined above, can be visually identified (by flickering the two images) a total of 34 changes between ${\bf X}$ and ${\bf Y}.$



Reference image (X)



5.

Test image (**Y**)











Ground truth (black regions denotes departures and white regions the arrivals) and ground truth with guard cells .

For each detector, the thresholds are set to ensure $P_{fa} = 10^{-3}$ in the complement of the extended ground truth area, namely, in the region where no changes occur;





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Performance Assessment



Unstructured and Structured GLRT





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Performance Assessment



Unstructured and Structured GLRT





	W		
Detector	3	5	7
Unstructured GLRT (6)	3802	6492	7533
Structured GLRT (8)	4949	6655	7387

Table: Number of correct detections for W = 3, 5 and 7.

- The Structured GLRT outperforms the unstructured GLRT for the smaller window sizes (W = 3 and 5) whereas the unstructured GLRT outperforms the structured GLRT for the larger window size of W = 7 when it is able to detect more changes in the image;
- This last result can be justified in terms of a covariance model mismatch in the sense that the off-diagonal entries of the polarimetric covariance matrix which in the theoretical model have been set to zero might not be exactly zero in reality (even if very close to that value).









Conclusions

- The block-diagonal structure for the polarimetric covariance matrix is exploited to derive a new decision rule based on the GLRT criterion;
- The proposed approach has been compared with both the optimal and the unstructured GLRT, with analysis on both simulated and real full-polarimetric SAR data;
- The performance analysis has confirmed that a structured approach can provide increase in performance with particular benefits when a small amount of homogeneous data is available;
- Possible future research tracks will consider the extension of the framework relaxing the Gaussian requirement for the data as well as the analysis on other datasets acquired by a different system, possibly at different resolutions (different performance behaviours could be observed on different datasets).









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